

CRAMER'S THEOREM FOR NONNEGATIVE MULTIVARIATE POINT PROCESSES WITH INDEPENDENT INCREMENTS

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ABSTRACT. We consider a continuous time version of Cramer's theorem with nonnegative summands $S_t = \frac{1}{t} \sum_{i:\tau_i \leq t} \xi_i$, $t \rightarrow \infty$, where $(\tau_i, \xi_i)_{i \geq 1}$ is a sequence of random variables such that tS_t is a random process with independent increments.

1. Introduction and main result

The following version of the Cramer theorem [1] can be extracted from Dembo and Zeitouni, [2].

Theorem 1. *Let $(\xi_i)_{i \geq 1}$ be a sequence of nonnegative identically distributed and independent random variables with ξ_1 admitting the Laplace transform:*

$$\mathcal{L}(\lambda) = \mathbb{E}e^{\lambda \xi_1}, \quad \lambda \in (-\infty, \Lambda), \quad \exists \Lambda = \inf\{\lambda > 0 : \mathcal{L}(\lambda) = \infty\}.$$

Then, the family

$$S_n = \frac{1}{n} \sum_{i=1}^n \xi_i, \quad n \rightarrow \infty$$

obeys the Large Deviation principle (LDP) in the metric space (\mathbb{R}_+, ρ) (for the Euclidean metric ρ) with the rate $\frac{1}{n}$ and the rate function

$$I(u) = \begin{cases} \sup_{\lambda \in (-\infty, \Lambda)} [\lambda u - g(\lambda)], & u > 0 \\ -\log \mathbb{P}(\xi_1 = 0), & u = 0, \end{cases}$$

where $g(\lambda)$ is the log moment generation function,

$$g(\lambda) = \log \mathbb{E}e^{\lambda \xi_1}, \quad \lambda < \Lambda.$$

In this paper, we study a “continuous time version” of this theorem. For $t \geq 0$, set

$$S_t = \frac{1}{t} \sum_{i:\tau_i \leq t} \xi_i,$$

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where $(\xi_i, \tau_i)_{i \geq 1}$ is a sequence of random pairs, where ξ_i 's and τ_i 's are random variables:

$$\xi_i \geq 0 \quad \text{and} \quad \tau_0 = 0 < \tau_1 < \tau_2 < \dots < \tau_i < \dots$$

defined on a probability space $(\Omega, \mathcal{F}, \mathbf{P})$. Let $(\mathcal{G}_n)_{n \geq 0}$ be the filtration with $\mathcal{G}_0 = (\emptyset, \Omega)$ and $\mathcal{G}_n := \sigma\{(\tau_i, \xi_i)_{i \leq n}\}$. Random variables ξ_i are assumed to be identically distributed and independent of \mathcal{G}_{i-1} with the distribution function

$$G(x) = \mathbf{P}(\xi_1 \leq x), \quad x \geq 0.$$

The conditional distribution of τ_i given \mathcal{G}_{i-1} is exponential:

$$\mathbf{P}(\tau_i \leq t | \mathcal{G}_{i-1}) = (1 - e^{-r(t-\tau_{i-1})}), \quad t \geq \tau_{i-1},$$

where r is a positive number. Moreover, we assume that

$$\mathbf{P}(\xi_i \leq x, \tau_i \leq t | \mathcal{G}_{i-1}) = G(x)(1 - e^{-r(t-\tau_{i-1})}), \quad i \geq 1.$$

The following theorem is an analogue of Theorem 1.

Theorem 2. *The family*

$$S_t = \frac{1}{t} \sum_{i: \tau_i \leq t} \xi_i, \quad t \rightarrow \infty$$

obeys the LDP in the metric space (\mathbb{R}_+, ϱ) with the rate $\frac{1}{t}$ and the rate function

$$I(u) = \begin{cases} \sup_{\lambda \in (-\infty, \Lambda)} [\lambda u - r \int_0^\infty (e^{\lambda z} - 1) dG(z)], & u > 0 \\ r[1 - G(0+)], & u = 0. \end{cases}$$

We give two examples illustrating compatibility with Theorems 1 and 2. For both discrete and continuous time cases, let

$$\mathbf{P}(\xi_1 \leq x) = 1 - e^{-x}, \quad x \geq 0,$$

so that, ξ_1 has the Laplace transform with $\Lambda = 1$. The log moment generating function is

$$g(\lambda) = -\log(1 - \lambda), \quad \lambda < 1.$$

For $G(x) = 1 - e^{-x}$, $x \geq 0$.

$$\int_0^\infty r(e^{\lambda z} - 1) dG(z) = \frac{r\lambda}{1 - \lambda}, \quad \lambda < 1.$$

In both cases, rate functions are explicitly computable (see also figures (1) and (2)),

$$I^d(u) = \begin{cases} u - 1 - \log(u), & u > 0 \\ \infty, & u = 0 \end{cases} \quad (\text{discrete time case})$$

$$I^c(u) = \begin{cases} (\sqrt{r} - \sqrt{u})^2, & u > 0 \\ r, & u = 0. \end{cases} \quad (\text{continuous time case})$$

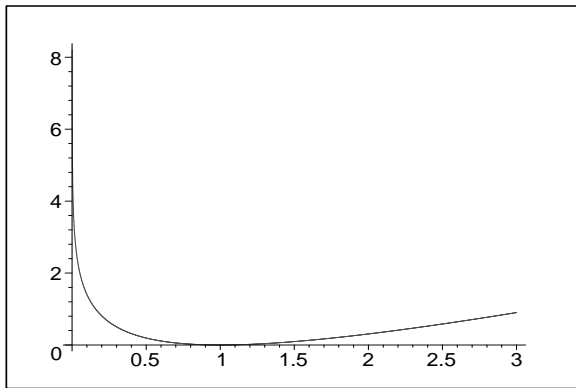


FIGURE 1. The rate function $I^d(u)$

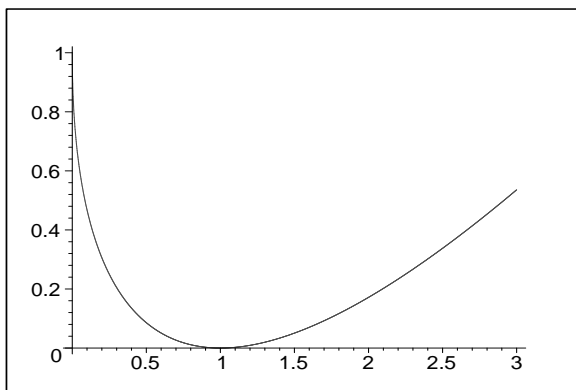


FIGURE 2. The rate function $I^c(u)$ for $r = 1$

Remark 1. Related topics to Theorem 2 can be found e.g. in Georgii and Zessin, [3], serving a class of marked point random fields. Probably, the proof of Theorem 2 can be adapted with arguments from proofs in [3] provided that many details not related to our setting have to be omitted and other ones concerning to the boundary effect have to be added.

We prefer to give a complete and direct proof of Theorem 2.

2. Counting random measure, its compensator, Laplace transform

We consider $(\xi_i, \tau_i)_{i \geq 1}$ as a multivariate (marked) point process (see, e.g. [4], [5]) with the counting measure

$$\mu(dt, dy) = \sum_{i \geq 1} \mathbf{I}_{\{\tau_i < \infty\}} \delta_{\{\tau_i, \xi_i\}}(t, y) dt dy,$$

where $\delta_{\{\tau_i, \xi_i\}}$ is the Dirac delta-function on $\mathbb{R}_+ \times \mathbb{R}_+$. Parallel to $(\mathcal{G}_n)_{n \geq 0}$, we introduce one more filtration $(\mathcal{G}_t)_{t \geq 0}$ related to $(\xi_i, \tau_i)_{i \geq 1}$:

$$\mathcal{G}_t := \sigma(\mu([0, t'] \times \Gamma) : t' \leq t, \Gamma \in \mathcal{B}(\mathbb{R}_+)),$$

where $\mathcal{B}(\mathbb{R}_+)$ is the Borel σ -algebra on \mathbb{R}_+ , and assume that \mathcal{G}_0 is augmented by \mathbb{P} -zero sets from \mathcal{F} (notice that, then, $(\mathcal{G}_t)_{t \geq 0}$ satisfies the general conditions). With the help of the counting measure μ , one can present tS_t in a form of a stochastic integral with respect to μ :

$$tS_t = \int_0^t \int_{x>0} x \mu(ds, dx). \quad (2.1)$$

Then, the Levy measure $\nu(ds, dx)$, related to μ , is explicitly computed (see, e.g. Theorem III.1.33, [5])

$$\nu(ds, dx) = \sum_{i \geq 1} I_{\llbracket \tau_{i-1}, \tau_i \rrbracket}(t) \frac{dG(x) de^{-r(s-\tau_{i-1})}}{e^{-r(s-\tau_{i-1})}} = r ds dG(x) \quad (2.2)$$

and is not random. It is well known (see, e.g. Corollary to Theorem 1 in §4, Ch.4, [6]) that under the deterministic Levy process tS_t is the random process with independent increments. We recall a useful property: for any nonnegative and (\mathcal{G}_t) -predictable function $f(\omega, x, t)$,

$$\mathbb{E} \int_0^t \int_{x>0} f(\omega, x, s) \mu(ds, dx) = \mathbb{E} \int_0^t \int_{x>0} f(\omega, x, s) \nu(ds, dx)$$

with “ $\infty = \infty$ ”.

Lemma 1. [Laplace transform] *For any $\lambda < \Lambda$ and $t > 0$,*

$$\mathbb{E} e^{\lambda t S_t} = e^{rt \int_{x>0} [e^{\lambda x} - 1] dG(x)}.$$

Proof. Though the direct computation of Laplace’s transform is permissible, we prefer to apply the stochastic calculus. The process $U_t = e^{\lambda t S_t}$ has right-continuous piece-wise constant paths with jumps

$$\Delta U_s = (U_s - U_{s-}) = U_{s-} \int_{x>0} [e^{\lambda x} - 1] \mu(\{s\}, dx),$$

so that, for any $t > 0$,

$$U_t = 1 + \int_0^t \int_{x>0} U_{s-} [e^{\lambda x} - 1] \mu(ds, dx).$$

Since the function $f(\omega, x, s) := U_{s-} [e^{\lambda x} + 1]$ is nonnegative and predictable, the following equality with $\lambda < \Lambda$ holds true:

$$\mathbb{E} \int_0^t \int_{x>0} U_{s-} [e^{\lambda x} + 1] \mu(ds, dx) = \mathbb{E} \int_0^t \int_{x>0} U_{s-} [e^{\lambda x} + 1] \nu(ds, dx) (< \infty).$$

Then, we also have

$$\mathbb{E} \int_0^t \int_{x>0} U_{s-} [e^{\lambda x} - 1] \mu(ds, dx) = \mathbb{E} \int_0^t \int_{x>0} U_{s-} [e^{\lambda x} - 1] \nu(ds, dx) (\in \mathbb{R}).$$

Since $\nu(ds, dx) = r ds dG(x)$, the later provides

$$(\mathbb{E} U_t) = 1 + \int_0^t \int_{x>0} (\mathbb{E} U_s) [e^{\lambda x} - 1] dG(x) r ds.$$

This can be written in an equivalent form of differential equation

$$\frac{d(\mathbb{E}U_t)}{dt} = (\mathbb{E}U_t) \int_{x>0} [e^{\lambda x} - 1] dG(x)r$$

subject to $(\mathbb{E}U_0) = 1$.

Thus, the desired result holds. \square

3. The proof of Theorem 2

We verify the necessary and sufficient conditions for the LDP to hold (for more details, see Puhalskii, [7]):

1) exponential tightness,

$$\lim_{j \rightarrow \infty} \overline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}(S_t \in \mathbb{R}_+ \setminus \mathcal{K}_j) = -\infty,$$

where \mathcal{K}_j 's are compacts increasing to \mathbb{R}_+ ;

2) local LDP, defining the rate function $I(u)$, $u \in \mathbb{R}_+$

$$\lim_{\delta \rightarrow 0} \lim_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}(|S_t - u| \leq \delta) = -I(u).$$

3.1. The exponential tightness. By choosing

$$K_j = \{x \in \mathbb{R}_+ : x \in [0, j]\}$$

and applying Chernoff's inequality with parameter 0.5Λ , we find that

$$\mathbb{P}(S_t > j) \leq e^{-0.5\Lambda j + \log \mathbb{E}e^{0.5\Lambda t S_t}}.$$

By Lemma 1,

$$\mathbb{E}e^{0.5\Lambda t S_t} = e^{rt \int_{x>0} [e^{0.5\Lambda x} - 1] dG(z)}$$

and, therefore,

$$\frac{1}{t} \log \mathbb{P}(S_t > j) \leq -0.5\Lambda j + r \int_{x>0} [e^{0.5\Lambda x} - 1] dG(x) \xrightarrow{j \rightarrow \infty} -\infty$$

and 1) is done.

3.2. The local LDP. We begin with computation of $I(0)$ and prove

$$\begin{aligned} \underline{\lim}_{\delta \rightarrow 0} \underline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}(S_t \leq \delta) &\geq -r[1 - G(0+)] \\ \overline{\lim}_{\delta \rightarrow 0} \overline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbb{P}(S_t \leq \delta) &\leq -r[1 - G(0+)] \end{aligned} \tag{3.1}$$

By (2.2), $\{tS_t = 0\} = \{\mu((0, t] \times \{x > 0\}) = 0\}$. Consequently for any $t > 0$,

$$\mathbb{P}(S_t \leq \delta) \geq \mathbb{P}(S_t = 0) = \mathbb{P}(tS_t = 0) = \mathbb{P}(\mu((0, t], \{x > 0\}) = 0).$$

The counting process $\pi_t := \mu((0, t], \{x > 0\})$ has independent increments and the rate

$$\mathbb{E}\mu((0, t], \{x > 0\}) = \nu((0, t], \{x > 0\}) = r[1 - G(0+)]t.$$

It is a counting process with the compensator $\nu((0, t], \{x > 0\}) = r[1 - G(0+)]t$. Therefore, by the Watanabe theorem, [8], π_t is a Poisson process with parameter $r[1 - G(0+)]$. Hence, due to well known property of the Poisson process

$$\mathbf{P}(\pi_t = 0) = e^{-tr[1-G(0+)]}.$$

We find that

$$\frac{1}{t} \log \mathbf{P}(S_t \leq \delta) \geq \frac{1}{t} \log \mathbf{P}(\pi_t = 0) = -r[1 - G(+)]$$

implying the lower bond from (3.1).

The upper bound from (3.1) is derived with the help of Laplace's transform with $0 < \lambda < \Lambda$. To this end, we use identity

$$1 = \mathbf{E} \exp \left(\lambda t S_t - tr \int_{x>0} [e^{\lambda x} - 1] dG(x) \right)$$

implying the inequality

$$1 \geq \mathbf{E} I_{\{S_t \leq \delta\}} \exp \left(t \left[\lambda \delta - r \int_{x>0} [e^{\lambda x} - 1] dG(x) \right] \right)$$

being equivalent to

$$\frac{1}{t} \log \mathbf{P}(S_t \leq \delta) \leq -\lambda \delta + r \int_{x>0} [e^{-\lambda x} - 1] dG(x).$$

Now, passing $t \rightarrow \infty$, we obtain the following upper bound depending on δ and λ :

$$\overline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbf{P}(S_t \leq \delta) \leq -\lambda \delta + r \int_{x>0} [e^{\lambda x} - 1] dG(x).$$

Now, passing $\delta \rightarrow 0$ and λ to $-\infty$ we find that

$$\overline{\lim}_{\delta \rightarrow 0} \overline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbf{P}(S_t \leq \delta) \leq -r \int_{x>0} dG(x) = -r[1 - G(0+)].$$

We continue the proof by checking the formula for $I(u)$ when $u > 0$, i.e.

$$\begin{aligned} \overline{\lim}_{\delta \rightarrow 0} \overline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbf{P}(|S_t - u| \leq \delta) &\leq -I(u) \\ \underline{\lim}_{\delta \rightarrow 0} \underline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbf{P}(|S_t - u| \leq \delta) &\geq -I(u), \end{aligned}$$

with

$$I(u) = \sup_{\lambda \in (-\infty, \Lambda)} \left[\lambda u - r \int_0^\infty (e^{\lambda x} - 1) dG(x) \right].$$

The Laplace transform

$$1 = \mathbf{E} \exp \left(\lambda t S_t - tr \int_{x>0} [e^{\lambda x} - 1] dG(x) \right), \quad \lambda < \Lambda,$$

implies the inequality

$$1 \geq \mathbf{E}I_{\{|S_t - u| \leq \delta\}} \exp \left(-t\delta u + \lambda t u - t r \int_{x>0} [e^{\lambda x} - 1] dG(x) \right)$$

providing the following upper bound depending on λ :

$$\overline{\lim}_{\delta \rightarrow 0} \overline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbf{P}(|S_t - u| \leq \delta) \leq - \left(\lambda u - r \int_{x>0} [e^{\lambda x} - 1] dG(x) \right).$$

A further minimization of the right hand side of the above inequality in λ over $(-\infty, \Lambda)$ gives the desired result.

The lower bound proof uses a standard approach of changing ‘‘probability measure’’. Denote by $\lambda^* = \operatorname{argmax}_{\lambda < \Lambda} (\lambda u - r \int_{x>0} [e^{\lambda x} - 1] dG(x))$. Since λ^* solves the equation (with $u > 0$)

$$u - r \int_{x>0} x e^{\lambda^* x} dG(x) = 0, \quad (3.2)$$

λ^* is a proper number strictly less than Λ . Set

$$\mathfrak{L}_t(\lambda^*) = \exp \left(\lambda^* t S_t - \int_0^t \int_{x>0} r [e^{\lambda^* x} - 1] dG(x) ds \right). \quad (3.3)$$

First of all we notice that the Laplace transform for tS_t with λ^* guarantees $\mathbf{E}\mathfrak{L}_t(\lambda^*) = 1$. Moreover, taking into account (2.1) and applying the Itô formula to $\mathfrak{L}_t(\lambda^*)$ one can see that $(\mathfrak{L}_t(\lambda^*), \mathcal{G}_t)_{t \geq 0}$ is a positive local martingale with paths from the Skorokhod space $\mathbb{D}_{[0, \infty)}$. Then a measure $\tilde{\mathbf{P}}_t$, defined by $d\tilde{\mathbf{P}}_t = \mathfrak{L}_t(\lambda^*) d\mathbf{P}_t$, where \mathbf{P}_t is a restriction of \mathbf{P} on \mathcal{G}_t , is the probability measure. We introduce the probability space $(\Omega, \mathcal{F}, \tilde{\mathbf{P}}_t)$. Since $\mathfrak{L}_t(\lambda^*) > 0$, \mathbf{P} -a.s., not only $\tilde{\mathbf{P}}_t \ll \mathbf{P}_t$ but also $\tilde{\mathbf{P}}_t \ll \mathbf{P}_t$ with $d\mathbf{P}_t = \mathfrak{L}_t^{-1}(\lambda^*) d\tilde{\mathbf{P}}_t$. This property and (3.3) provide a lower bound

$$\begin{aligned} \mathbf{P}(|S_t - u| \leq \delta) &= \mathbf{P}_t(|S_t - u| \leq \delta) = \int_{\{|S_t - u| \leq \delta\}} \mathfrak{L}_t^{-1}(\lambda^*) d\tilde{\mathbf{P}}_t \\ &\geq e^{-\lambda^* t \delta - t I(u)} \tilde{\mathbf{P}}_t(|S_t - u| \leq \delta). \end{aligned}$$

Therefore, we find that

$$\underline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \mathbf{P}(|S_t - u| \leq \delta) \geq -I(u) - \lambda^* \delta + \underline{\lim}_{t \rightarrow \infty} \frac{1}{t} \log \tilde{\mathbf{P}}_t(|S_t - u| \leq \delta).$$

Obviously, the desired lower bound to obtain it is left to prove that

$$\underline{\lim}_{t \rightarrow \infty} \tilde{\mathbf{P}}_t(|S_t - u| \leq \delta) = 1$$

or, equivalently,

$$\lim_{t \rightarrow \infty} \tilde{\mathbf{P}}_t(|S_t - u| > \delta) = 0. \quad (3.4)$$

Thus the last step of the proof deal with (3.4). To this end, we show that $(\tilde{\mathbf{E}}_t$ denotes the expectation relative to $\tilde{\mathbf{P}}_t)$

$$\lim_{t \rightarrow \infty} \tilde{\mathbf{E}}_t |S_t - u|^2 = 0.$$

Since $\mathbf{E}\mathfrak{L}_t(\lambda) = 1$, it holds

$$\begin{aligned} 0 &= \frac{\partial^2 \mathbf{E}\mathfrak{L}_t(\lambda)}{\partial \lambda^2} \Big|_{\lambda=\lambda^*} \\ &= t^2 \mathbf{E} \left(S_t - r \int_{\{x>0\}} x e^{\lambda^* x} dG(x) \right)^2 \mathfrak{L}_t(\lambda^*) - t r \underbrace{\int_{\{x>0\}} x^2 e^{\lambda^* x} dG(x)}_{=u \text{ (see (3.2))}} \\ &= t^2 \tilde{\mathbf{E}} \left(S_t - r \int_{\{x>0\}} x e^{\lambda^* x} dG(x) \right)^2 - tu. \end{aligned}$$

Hence,

$$\tilde{\mathbf{E}}(S_t - u)^2 = \frac{1}{t} r \mathbf{E} \int_{\{x>0\}} x^2 e^{\lambda^* x} dG(x) \xrightarrow{t \rightarrow \infty} 0.$$

□

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